

# Asset Pricing Models

Econ 643: Financial Economics II

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- Unifying approach to pricing stocks, bonds and derivative products
  - (1) is referred to as fundamental pricing equation (Cochrane, 2001).

# Stochastic Discount Factor Approach

- Dividing both sides of (1) by  $p_t$  (assuming non-zero prices) and rearranging

$$E_t(m_{t+1}R_{t+1} - 1) = 0$$

or equivalently

$$E(m_{t+1}R_{t+1} - 1 | I_t) = 0$$

where  $R_{t+1} = \frac{x_{t+1}}{p_t} = \frac{p_{t+1} + d_{t+1}}{p_t}$  denotes the gross asset return.

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- Portfolios based on excess returns  $R^e = R - R^f$ , where  $R^f$  denotes the risk-free rate, are zero-cost portfolios (borrow one dollar at interest rate  $R^f$  and invest it in a asset with return  $R$ ).

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- In this case, with zero price and payoff  $R_{t+1}^e = R_{t+1} - R_{t+1}^f$ , the fundamental pricing equation is given by

$$E_t(m_{t+1}R_{t+1}^e) = 0.$$

# Stochastic Discount Factor Approach: An Example

- Suppose that a representative agent maximizes expected utility

$$\sum_{t=1}^{\infty} \beta^t E[u(c_t) | I_0]$$

subject to a budget constraint

$$a_{t+1} = (a_t + y_t - c_t) R_{t+1},$$

where  $c_t$  is consumption,  $a_t$  is an asset with gross return  $R_t$ ,  $y_t$  is income and  $I_t$  is the information set at time  $t$ .

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$$E \left[ \beta \frac{u'(c_{t+1})}{u'(c_t)} R_{t+1} - 1 \mid I_t \right] = 0.$$

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- It takes the form of the fundamental pricing equation with the SDF  $m_{t+1} = \beta \frac{u'(c_{t+1})}{u'(c_t)}$  given by the intertemporal marginal rate of substitution.

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- No arbitrage  $\Leftrightarrow$  positive SDF; negative SDF  $\not\Rightarrow$  arbitrage opportunity.
- But imposing explicitly the no-arbitrage constraint may have some adverse effects.
- More on how and whether we should impose positivity of the SDF, see Gospodinov, Kan and Robotti (2010).

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$$\begin{aligned}E(R^i) &= \frac{1}{E(m)} - \frac{\text{Cov}(m, R^i)}{E(m)} \\ &= \frac{1}{E(m)} + \left[ \frac{\text{Cov}(m, R^i)}{\text{Var}(m)} \right] \left[ -\frac{\text{Var}(m)}{E(m)} \right] \\ &= \gamma + \beta_{i,m}\lambda_m\end{aligned}$$

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using that  $1/E(m) = R^f = \gamma$  from above.

- $\beta_{i,m} = \text{Cov}(m, R^i) / \text{Var}(m)$  is the regression coefficient of the return  $R^i$  on  $m$  and  $\lambda_m = -\text{Var}(m) / E(m) < 0$  is the price of risk.

# Beta Representation

- Some interesting observations emerge from rewriting

$$E(R^i) = \frac{1}{E(m)} - \frac{\text{Cov}(m, R^i)}{E(m)}$$

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  - therefore, assets that have returns positively correlated with consumption should pay more than the risk-free rate.
- If  $\text{Cov}(c, R^i) < 0$ , then the expected excess returns are negative.
  - these assets provide insurance against bad outcome and smooth consumption.

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- Then, substituting into the fundamental pricing equation and rearranging (see Cochrane, 2005, pp.107-108), we get

$$E(R^i) = \gamma + \lambda' \beta_i,$$

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- $\lambda_k$  is the price of this risk exposure.

# Factor Models

- Let  $i = 1, 2, \dots, N$  be the number of assets,  $j = 1, 2, \dots, k$  be the number of risk factors and  $t = 1, 2, \dots, T$  denote the number of time series observations on these assets and factors.

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- First,  $\beta$ 's are estimated from the *time series regression* for each asset  $i$

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- if there is a risk-free asset,  $\gamma$  is the return on this asset

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$$E(R^i) = \gamma + \beta_{i,1}\lambda_1 + \dots + \beta_{i,k}\lambda_k + \alpha_i, \quad i = 1, 2, \dots, N.$$

- $\alpha_i$  are *pricing errors*; model predicts  $\alpha_i = 0$
- $\gamma$  is the expected zero-beta rate, i.e. the expected return of any security that is uncorrelated with each of the factors ( $\beta_{0,j} = 0$  for all  $j$ )
- if there is a risk-free asset,  $\gamma$  is the return on this asset
- $\gamma$  and  $\lambda$  are the same for all assets.

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- If the factors are non-traded factors (e.g., macroeconomic factors),  $\lambda$  is the model's predicted price rather than a market price of the factor.
- **A test of  $\lambda = 0$  is a test of whether the factor is priced or not.**

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- The model implies  $E(R^{ei}) = \beta_i E(f)$  and the pricing errors should be jointly equal to zero.

# Evaluation of Factor Models

- To test if the model is correctly specified (i.e. the pricing errors are zero), use the statistic

$$\hat{\alpha}' [\text{Var}(\hat{\alpha})]^{-1} \hat{\alpha} = T \left[ 1 + \left( \frac{\hat{\mu}_f}{\hat{\sigma}_f} \right)^2 \right]^{-1} \hat{\alpha}' \hat{\Sigma}^{-1} \hat{\alpha} \sim \chi_N^2,$$

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- If the pricing errors are large, the test statistic will exceed the critical value from the chi-square distribution with  $N$  degrees of freedom and the null hypothesis will be rejected.

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- 2 Second pass (cross-sectional regression): estimating the factor risk premia  $\lambda$  and pricing errors  $\alpha$  from the equation

$$\bar{R}^{ei} = \hat{\beta}_i' \lambda + \alpha_i, \quad i = 1, 2, \dots, N$$

as  $\hat{\lambda} = (\hat{\beta}'\hat{\beta})^{-1}\hat{\beta}'\bar{R}^e$  (OLS) or  $\hat{\lambda} = (\hat{\beta}'\hat{\Sigma}^{-1}\hat{\beta})^{-1}\hat{\beta}'\hat{\Sigma}^{-1}\bar{R}^e$  (GLS) and  $\hat{\alpha} = \bar{R}^e - \hat{\lambda}'\hat{\beta}$ , where  $\bar{R}^{ei} = \frac{1}{T} \sum_{t=1}^T R_t^{ei}$ .

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  - hence, the conventional standard errors should be adjusted to reflect the estimation error in the second-pass regressors  $\hat{\beta}$
  - we will see later that this correction does not need to be performed explicitly if the model is estimated by GMM.

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- In practice, this is not the case.
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- We will consider first the statistical approach based on factor analysis and principal components.

# Statistical Selection of Factors

- Suppose that we have access to a large panel of data  $x_{it}$  ( $i=1, \dots, M$ ,  $t=1, \dots, T$ ), where  $M$  is the number of variables (returns, macro variables) and  $T$  is the number of time series observations.

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- Concentrating out  $(\omega_1', \dots, \omega_M')'$ , the problem of estimating  $f_t$  is identical to maximizing  $tr(F'(X'X)F)$  and  $\tilde{f}_t$  is  $\sqrt{T}$  times the  $k$  eigenvectors corresponding to the  $k$  largest eigenvalues of the matrix  $XX'/(MT)$ .

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- Fama-French (1993) five-factor model
  - FF three-factor model plus term structure factor (difference in the yields on 30-year bond and 1-month bill) and default premium (difference between the yields on BAA and AAA corporate bonds).

# Conditional Asset Pricing Models

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- The conditional asset pricing model that satisfies  $p_t = E_t(m_{t+1}x_{t+1})$  is the model

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where  $a_t$  and  $b_t$  are possibly time-varying.

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- In a model with one factor and one conditioning variable

$$\begin{aligned}m_{t+1} &= (a_0 + a_1 z_t) + (b_0 + b_1 z_t) f_{t+1} \\ &= a_0 + a_1 z_t + b_0 f_{t+1} + b_1 (z_t f_{t+1}).\end{aligned}$$

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- therefore, we can use the new (scaled) factors with the unconditional moment procedure that we developed above.
- When we have many factors and instruments, we obtain the scaled factors by multiplying each factor by each instrument using the Kronecker product

$$\text{scaled factors} = f_{t+1} \otimes z_t.$$

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  - **stock return volatility: realized or implied volatility.**

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    - compare its performance (using Sharpe ratio) to a buy-and-hold benchmark strategy over the out-of-sample evaluation period.